

Archimax Copulas

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<http://freakonometrics.hypotheses.org/>

based on joint work with

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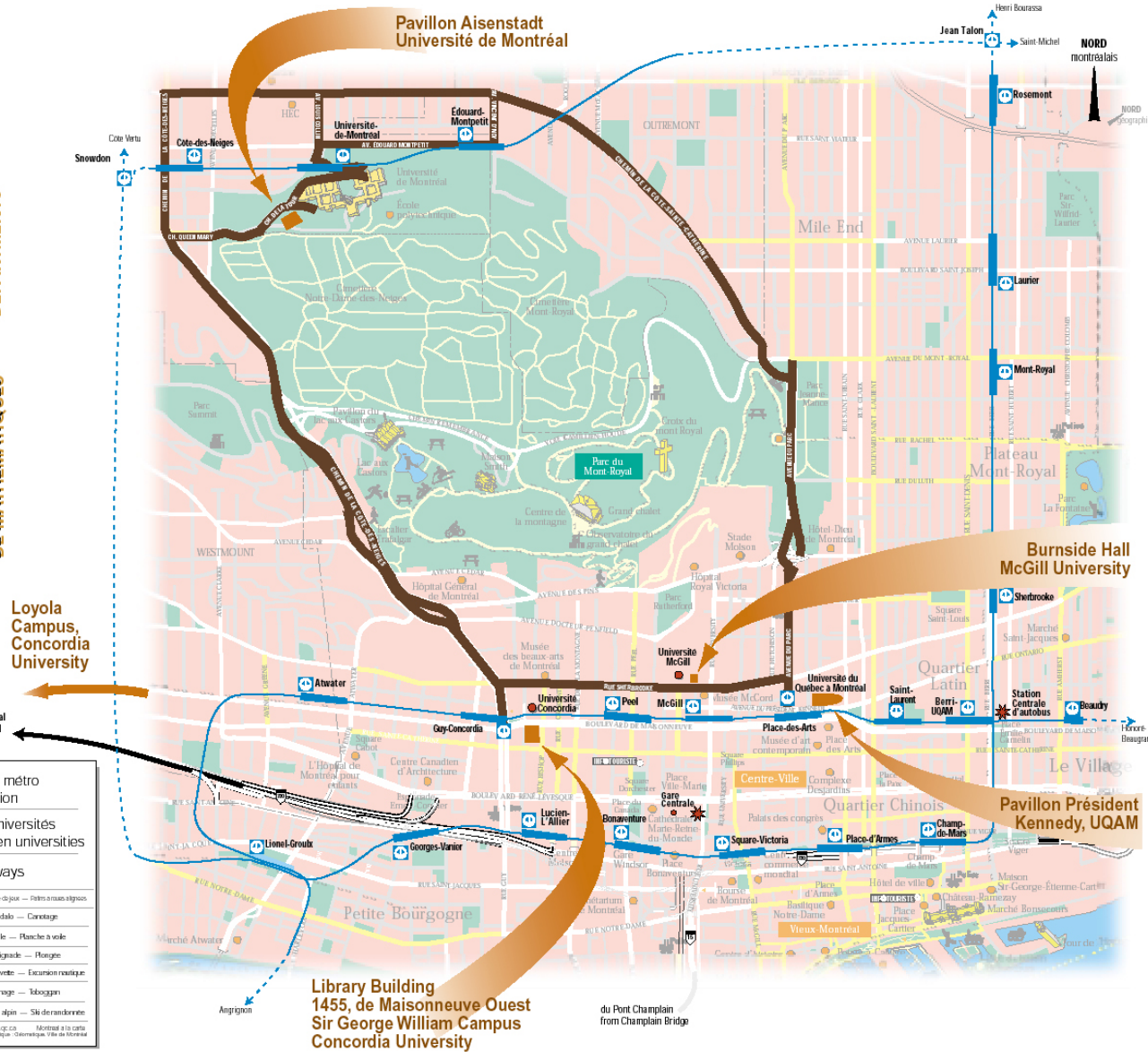


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MONTRÉAL

AND ITS FOUR MATHEMATICS DEPARTMENTS

ET SES QUATRE DÉPARTEMENTS DE MATHÉMATIQUES



Agenda

- Copulas
- **Standard copula families**
- Elliptical distributions (and copulas)
- Archimedean copulas
- Extreme value distributions (and copulas)
- **Archimax copulas**
- Archimax copulas in dimension 2
- Archimax copulas in dimension $d \geq 3$

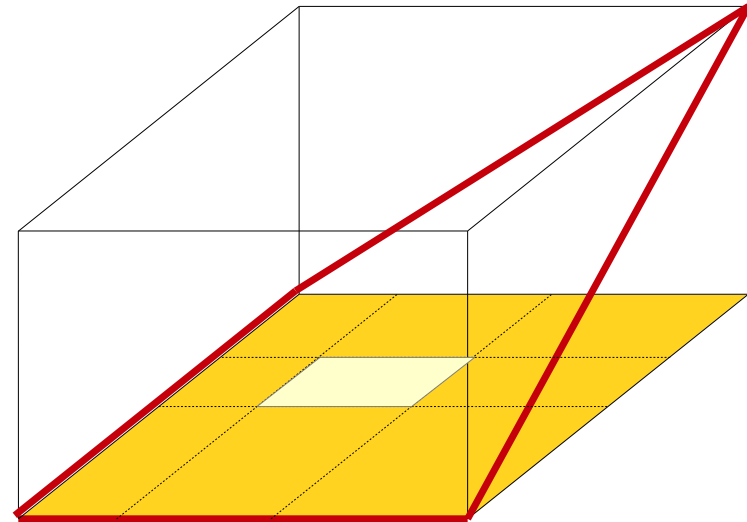
Copulas, in dimension $d = 2$

Definition 1

A **copula** in dimension 2 is a c.d.f on $[0, 1]^2$, with margins $\mathcal{U}([0, 1])$.

Thus, let $C(u, v) = \mathbb{P}(U \leq u, V \leq v)$,
where $0 \leq u, v \leq 1$, then

- $C(0, x) = C(x, 0) = 0 \quad \forall x \in [0, 1]$,
- $C(1, x) = C(x, 1) = x \quad \forall x \in [0, 1]$,
- and some *increasingness* property



Copulas, in dimension $d = 2$

Definition 2

A **copula** in dimension 2 is a c.d.f on $[0, 1]^2$, with margins $\mathcal{U}([0, 1])$.

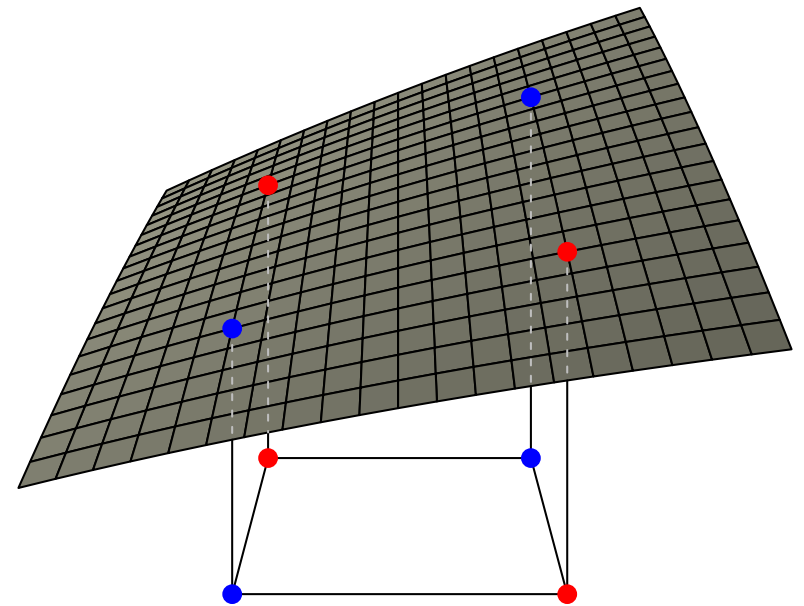
Thus, let $C(u, v) = \mathbb{P}(U \leq u, V \leq v)$,
where $0 \leq u, v \leq 1$, then

- $C(0, x) = C(x, 0) = 0 \quad \forall x \in [0, 1]$,
- $C(1, x) = C(x, 1) = x \quad \forall x \in [0, 1]$,
- If $0 \leq u_1 \leq u_2 \leq 1, 0 \leq v_1 \leq v_2 \leq 1$

$$C(u_2, v_2) + C(u_1, v_1) \geq C(u_1, v_2) + C(u_2, v_1)$$

(concept of 2-increasing function in \mathbb{R}^2)

see $C(u, v) = \int_0^v \int_0^u \underbrace{c(x, y)}_{\geq 0} dx dy$ with the density notation.



Copulas, in dimension $d \geq 2$

The concept of d -increasing function simply means that

$$\mathbb{P}(a_1 \leq U_1 \leq b_1, \dots, a_d \leq U_d \leq b_d) = \mathbb{P}(\mathbf{U} \in [\mathbf{a}, \mathbf{b}]) \geq 0$$

where $\mathbf{U} = (U_1, \dots, U_d) \sim C$ for all $\mathbf{a} \leq \mathbf{b}$ (where $a_i \leq b_i$).

Definition 3

Function $h : \mathbb{R}^d \rightarrow \mathbb{R}$ is d -increasing if for all rectangle $[\mathbf{a}, \mathbf{b}] \subset \mathbb{R}^d$, $V_h([\mathbf{a}, \mathbf{b}]) \geq 0$, where

$$V_h([\mathbf{a}, \mathbf{b}]) = \Delta_{\mathbf{a}}^{\mathbf{b}} h(\mathbf{t}) = \Delta_{a_d}^{b_d} \Delta_{a_{d-1}}^{b_{d-1}} \dots \Delta_{a_2}^{b_2} \Delta_{a_1}^{b_1} h(\mathbf{t}) \quad (1)$$

and for all \mathbf{t} , with

$$\Delta_{a_i}^{b_i} h(\mathbf{t}) = h(t_1, \dots, t_{i-1}, b_i, t_{i+1}, \dots, t_n) - h(t_1, \dots, t_{i-1}, a_i, t_{i+1}, \dots, t_n). \quad (2)$$

Copulas, in dimension $d \geq 2$

Definition 4

A **copula** in dimension d is a c.d.f on $[0, 1]^d$, with margins $\mathcal{U}([0, 1])$.

Theorem 1 1. If C is a copula, and F_1, \dots, F_d are univariate c.d.f., then

$$F(x_1, \dots, x_n) = C(F_1(x_1), \dots, F_d(x_d)) \quad \forall (x_1, \dots, x_d) \in \mathbb{R}^d \quad (3)$$

is a multivariate c.d.f. with $F \in \mathcal{F}(F_1, \dots, F_d)$.

2. Conversely, if $F \in \mathcal{F}(F_1, \dots, F_d)$, there exists a copula C satisfying (3). This copula is usually not unique, but it is if F_1, \dots, F_d are absolutely continuous, and then,

$$C(u_1, \dots, u_d) = F(F_1^{-1}(u_1), \dots, F_d^{-1}(u_d)), \quad \forall (u_1, \dots, u_d) \in [0, 1]^d \quad (4)$$

where quantile functions $F_1^{-1}, \dots, F_n^{-1}$ are generalized inverse (left cont.) of F_i 's.

If $\mathbf{X} \sim F$, then $\mathbf{U} = (F_1(X_1), \dots, F_d(X_d)) \sim C$.

Survival (or dual) copulas

Theorem 2 1. If C^* is a copula, and $\bar{F}_1, \dots, \bar{F}_d$ are univariate s.d.f., then

$$\bar{F}(x_1, \dots, x_n) = C^*(\bar{F}_1(x_1), \dots, \bar{F}_d(x_d)) \quad \forall (x_1, \dots, x_d) \in \mathbb{R}^d \quad (5)$$

is a multivariate s.d.f. with $F \in \mathcal{F}(F_1, \dots, F_d)$.

2. Conversely, if $F \in \mathcal{F}(F_1, \dots, F_d)$, there exists a copula C^* satisfying (5). This copula is usually not unique, but it is if F_1, \dots, F_d are absolutely continuous, and then,

$$C^*(u_1, \dots, u_d) = \bar{F}(\bar{F}_1^{-1}(u_1), \dots, \bar{F}_d^{-1}(u_d)), \quad \forall (u_1, \dots, u_d) \in [0, 1]^d \quad (6)$$

where quantile functions $F_1^{-1}, \dots, F_n^{-1}$ are generalized inverse (left cont.) of F_i 's.

If $\mathbf{X} \sim F$, then $\mathbf{U} = (F_1(X_1), \dots, F_d(X_d)) \sim C$ and $\mathbf{1} - \mathbf{U} \sim C^*$.

Benchmark copulas

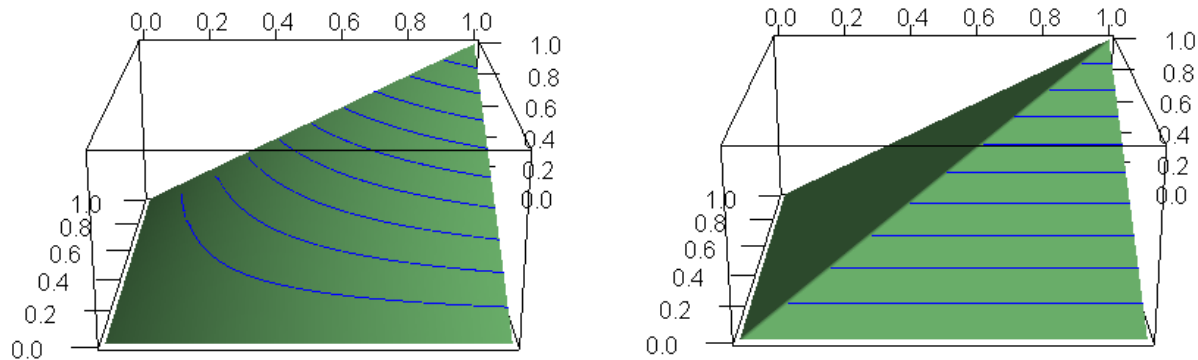
Definition 5

The independent copula C^\perp is defined as

$$C^\perp(u_1, \dots, u_d) = u_1 \times \dots \times u_d = \prod_{i=1}^d u_i.$$

Definition 6

The comonotonic copula C^+ (the Fréchet-Hoeffding upper bound of the set of copulas) is the copula defined as $C^+(u_1, \dots, u_d) = \min\{u_1, \dots, u_d\}$.



Spherical distributions

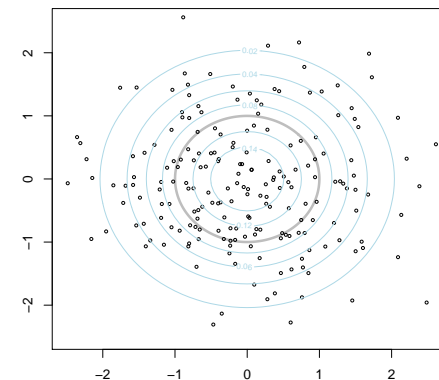
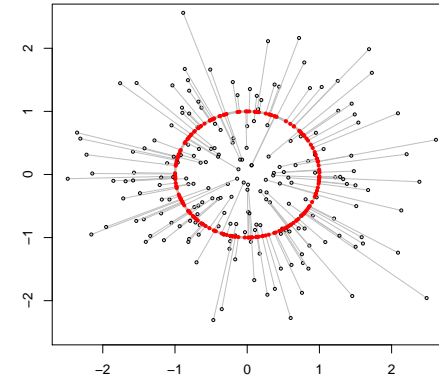
Definition 7

Random vector \mathbf{X} as a spherical distribution if

$$\mathbf{X} = R \cdot \mathbf{U}$$

where R is a positive random variable and \mathbf{U} is uniformly distributed on the unit sphere of \mathbb{R}^d , with $R \perp\!\!\!\perp \mathbf{U}$.

E.g. $\mathbf{X} \sim \mathcal{N}(\mathbf{0}, \mathbb{I})$.



Those distribution can be non-symmetric, see Hartman & Wintner (AJM, 1940) or Cambanis, Huang & Simons (JMVA, 1979))

Elliptical distributions

Definition 8

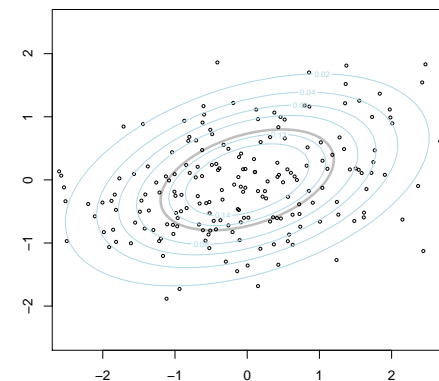
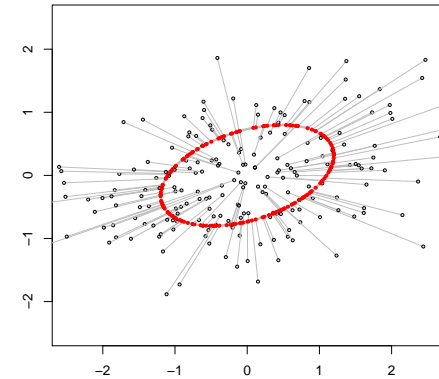
Random vector \mathbf{X} as a **elliptical distribution** if

$$\mathbf{X} = \boldsymbol{\mu} + R \cdot \mathbf{A} \cdot \mathbf{U}$$

where R is a positive random variable and \mathbf{U} is uniformly distributed on the unit sphere of \mathbb{R}^d , with $R \perp \mathbf{U}$, and where \mathbf{A} satisfies $\mathbf{A}\mathbf{A}' = \boldsymbol{\Sigma}$.

E.g. $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.

Elliptical distribution are popular in finance, see e.g. Jondeau, Poon & Rockinger (FMPM, 2008)



Archimedean copula

Definition 9

If $d \geq 2$, an **Archimedean generator** is a function $\phi : [0, 1] \rightarrow [0, \infty)$ such that ϕ^{-1} is d -completely monotone (i.e. ψ is d -completely monotone if ψ is continuous and $\forall k = 0, 1, \dots, d, (-1)^k d^k \psi(t) / dt^k \geq 0$).

Definition 10

Copula C is an **Archimedean copula** is, for some generator ϕ ,

$$C(u_1, \dots, u_d) = \phi^{-1}[\phi(u_1) + \dots + \phi(u_d)], \forall u_1, \dots, u_d \in [0, 1].$$

Exemple 1

$\phi(t) = -\log(t)$ yields the independent copula C^\perp .

$\phi(t) = [-\log(t)]^\theta$ yields Gumbel copula C_θ (note that $\psi(t) = \phi^{-1}(t) = \exp[-t^{1/\theta}]$).

Archimedean copula, exchangeability and frailties

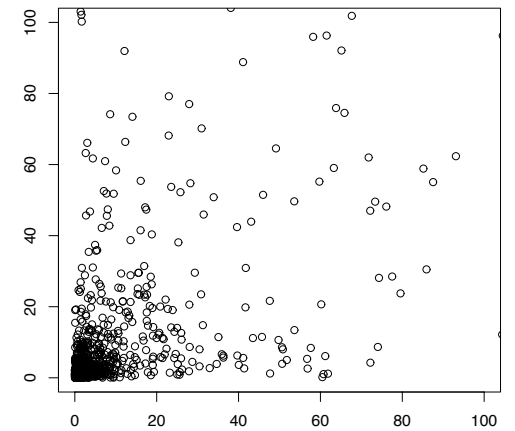
Consider residual life times $\mathbf{X} = (X_1, \dots, X_d)$ **conditionally independent** given some latent factor Θ , and such that $\mathbb{P}(X_i > x_i | \Theta = \theta) = \bar{B}_i(x_i)^\theta$. Then

$$\bar{F}(\mathbf{x}) = \mathbb{P}(\mathbf{X} > \mathbf{x}) = \psi \left(- \sum_{i=1}^n \log \bar{F}_i(x_i) \right)$$

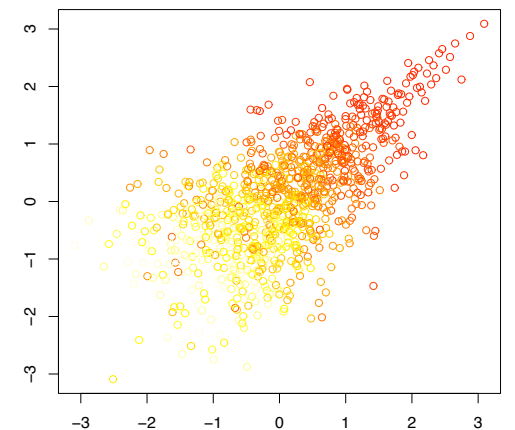
where ψ is the Laplace transform of Θ , $\psi(t) = \mathbb{E}(e^{-t\Theta})$. Thus, the survival copula of \mathbf{X} is Archimedean, with generator $\phi = \psi^{-1}$.

See Oakes (JASA, 1989).

Conditional independence, continuous risk factor



Conditional independence, continuous risk factor



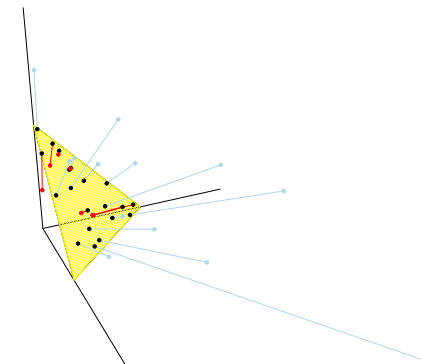
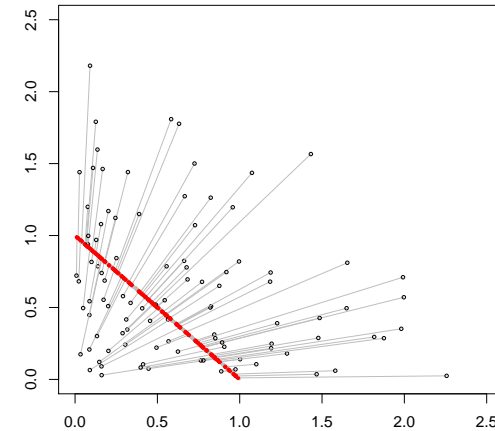
Stochastic representation of Archimedean copulas

Consider some strictly positive random variable R independent of \mathbf{U} , uniform on the simplex of \mathbb{R}^d . The survival copula of $\mathbf{X} = R \cdot \mathbf{U}$ is Archimedean, and its generator is the inverse of Williamson d -transform,

$$\phi^{-1}(t) = \int_x^\infty \left(1 - \frac{x}{t}\right)^{d-1} dF_R(t).$$

Note that $R \stackrel{\mathcal{L}}{=} \phi(U_1) + \dots + \phi(U_d)$.

See Nešlehová & McNeil (AS, 2009).



Archimedean copula and distortion

Definition 11

Function $h : [0, 1] \rightarrow [0, 1]$ defined as $h(t) = \exp[-\phi(t)]$ is called a **distortion function**.

Genest & Rivest (SPL, 2001), Morillas (M, 2005) considered distorted copulas (also called *multivariate probability integral transformation*)

Definition 12

Let h be some distortion function, and C a copula, then

$$C_h(u_1, \dots, u_d) = h^{-1}(C(h(u_1), \dots, h(u_d)))$$

is a copula.

Exemple2

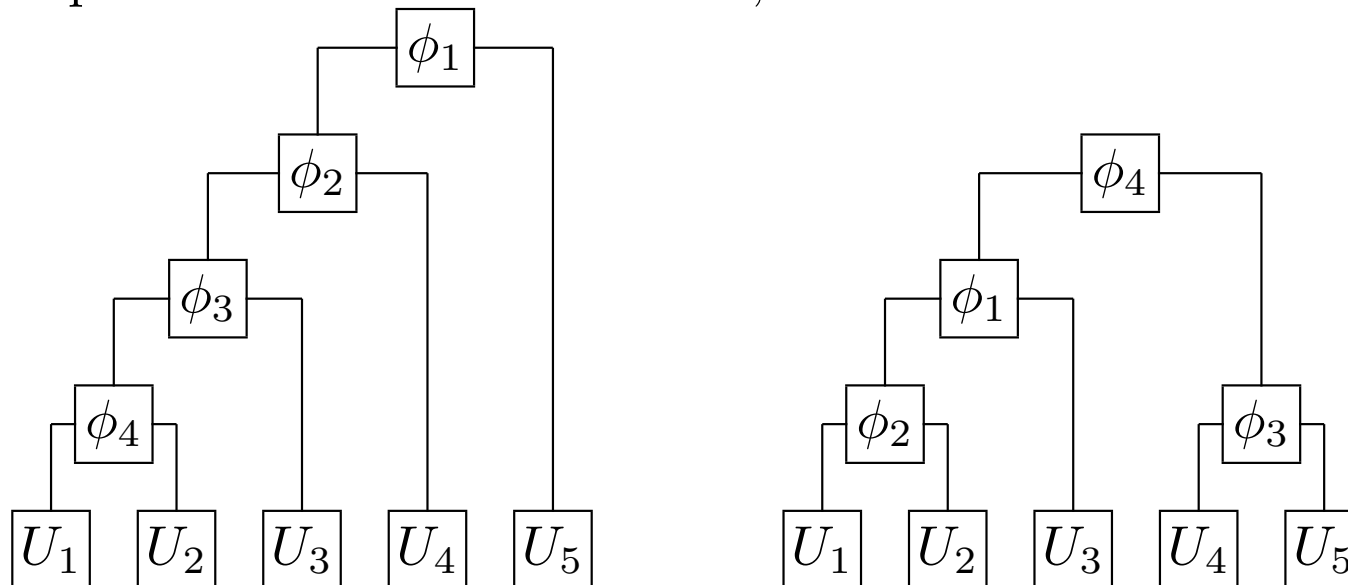
If $C = C^\perp$, then C_h^\perp is the Archimedean copula with generator $\phi(t) = -\log h(t)$.

Nested Archimedean copula, and hierarchical structures

Consider $C(u_1, \dots, u_d)$ defined as

$$\phi_1^{-1}[\phi_1[\phi_2^{-1}(\phi_2[\dots\phi_{d-1}^{-1}[\phi_{d-1}(u_1) + \phi_{d-1}(u_2)] + \dots + \phi_2(u_{d-1}))] + \phi_1(u_d)]]$$

where ϕ_i 's are generators. Then C is a copula if $\phi_i \circ \phi_{i-1}^{-1}$ is the inverse of a Laplace transform, and is called **fully nested Archimedean copula**. Note that partial nested copulas can also be considered,



(Univariate) extreme value distributions

Central limit theorem, $X_i \sim F$ i.i.d. $\frac{\bar{X}_n - b_n}{a_n} \xrightarrow{\mathcal{L}} S$ as $n \rightarrow \infty$ where S is a non-degenerate random variable.

Fisher-Tippett theorem, $X_i \sim F$ i.i.d., $\frac{X_{n:n} - b_n}{a_n} \xrightarrow{\mathcal{L}} M$ as $n \rightarrow \infty$ where M is a non-degenerate random variable.

Then

$$\mathbb{P} \left(\frac{X_{n:n} - b_n}{a_n} \leq x \right) = F^n(a_n x + b_n) \rightarrow G(x) \text{ as } n \rightarrow \infty, \forall x \in \mathbb{R}$$

i.e. F belongs to the max domain of attraction of G , G being an **extreme value distribution** : the limiting distribution of the normalized maxima.

$$-\log G(x) = (1 + \xi x)_+^{-1/\xi}$$

(Multivariate) extreme value distributions

Assume that $\mathbf{X}_i \sim F$ i.i.d.,

$$F^n(\mathbf{a}_n \mathbf{x} + \mathbf{b}_n) \rightarrow G(\mathbf{x}) \text{ as } n \rightarrow \infty, \forall \mathbf{x} \in \mathbb{R}^d$$

i.e. F belongs to the max domain of attraction of G , G being an (multivariate) extreme value distribution : the limiting distribution of the normalized componentwise maxima,

$$\mathbf{X}_{n:n} = (\max\{X_{1,i}\}, \dots, \max\{X_{d,i}\})$$

$$-\log G(\mathbf{x}) = \mu([\mathbf{0}, \infty) \setminus [\mathbf{0}, \mathbf{x}]), \forall \mathbf{x} \in \mathbb{R}_+^d$$

where μ is the exponent measure. It is more common to use the stable tail dependence function ℓ defined as

$$\ell(\mathbf{x}) = \mu([\mathbf{0}, \infty) \setminus [\mathbf{0}, \mathbf{x}^{-1}]), \forall \mathbf{x} \in \mathbb{R}_+^d$$

i.e.

$$-\log G(\mathbf{x}) = \ell(-\log G_1(x_1), \dots, \log G_d(x_d)), \forall \mathbf{x} \in \mathbb{R}^d$$

Note that there exists a finite measure H on the simplex of \mathbb{R}^d such that

$$\ell(x_1, \dots, x_d) = \int_{\mathcal{S}_d} \max\{\omega_1 x_1, \dots, \omega_d x_d\} dH(\omega_1, \dots, \omega_d)$$

for all $(x_1, \dots, x_d) \in \mathbb{R}_+^d$, and $\int_{\mathcal{S}_d} \omega_i dH(\omega_1, \dots, \omega_d) = 1$ for all $i = 1, \dots, n$.

Definition 13

Copula $C : [0, 1]^d \rightarrow [0, 1]$ is an **multivariate extreme value copula** if and only if there exists a **stable tail dependence function** such that ℓ

$$C_\ell(u_1, \dots, u_d) = \exp[-\ell(-\log u_1, \dots, -\log u_d)]$$

Assume that $U_i \sim \Gamma$ i.i.d.,

$$\Gamma^n(\mathbf{u}^{\frac{1}{n}}) = \Gamma^n(u_1^{\frac{1}{n}}, \dots, u_d^{\frac{1}{n}}) \rightarrow C_\ell(\mathbf{u}) \text{ as } n \rightarrow \infty, \forall \mathbf{x} \in \mathbb{R}^d$$

i.e. Γ belongs to the max domain of attraction of C_ℓ , C_ℓ being an **(multivariate) extreme value copula**, $\Gamma \in \text{MDA}(C_\ell)$.

The stable tail dependence function $\ell(\cdot)$

Observe that

$$n \left[1 - C \left(1 - \frac{x_1}{n}, \dots, 1 - \frac{x_d}{n} \right) \right] \rightarrow \underbrace{-\log [\Gamma(e^{-x_1}, \dots, e^{-x_d})]}_{=\ell(\mathbf{x})}$$

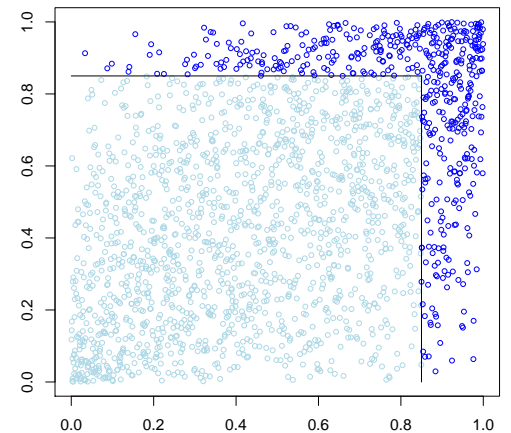
Exemple3

Gumbel copula, $\theta \in [1, +\infty]$,

$$\ell_\theta(x_1, \dots, x_d) = (x_1^\theta + \dots + x_d^\theta)^{1/\theta} = \|\mathbf{x}\|_\theta \quad \forall \mathbf{x} \in \mathbb{R}_+^d$$

Function $\ell(\cdot)$ satisfies

$$\underbrace{\max\{x_1, \dots, x_d\}}_{\substack{\text{(asympt.) comonotonicity} \\ \ell_\infty(\mathbf{x})}} \leq \ell(\mathbf{x}) \leq \underbrace{x_1 + \dots + x_d}_{\substack{\text{(asympt.) independence} \\ \ell_1(\mathbf{x})}}$$



The stable tail dependence function $\ell(\cdot)$

Function $\ell(\cdot)$ is homogeneous, $\ell(t \cdot \mathbf{x}) = t \cdot \ell(\mathbf{x}) \forall t \in \mathbb{R}_+$.

→ consider the restriction of $\ell(\cdot)$ on the unit simplex Δ_{d-1} ,

$$\ell(\mathbf{x}) = \|\mathbf{x}\|_1 \cdot \underbrace{\ell\left(\frac{x_1}{\|\mathbf{x}\|}, \dots, \frac{x_d}{\|\mathbf{x}\|}\right)}_{\ell(\boldsymbol{\omega})} = \|\mathbf{x}\|_1 \cdot A(\omega_1, \dots, \omega_{d-1})$$

where $A(\cdot)$ is **Pickands dependence function**. Observe that

$$\max\{\omega_1, \dots, \omega_{d-1}, \omega_d\} \leq A(\omega_1, \dots, \omega_{d-1}) \leq 1, \quad \forall \boldsymbol{\omega} \in \Delta_{d-1}$$

What do we have in dimension 2 ?

C is an Archimedean copula if $C = C_\phi$

$$C_\phi(u, v) = \phi^{-1} [\phi(u) + \phi(v)]$$

C is an extreme value copula if $C = C_A = C_\ell$

$$\begin{cases} C_A(u, v) = \exp \left(\log[uv] A \left(\frac{\log[v]}{\log[uv]} \right) \right) \\ C_\ell(u, v) = \exp[-\ell(-\log u, -\log v)] \end{cases}$$

where $A : [0, 1] \rightarrow [1/2, 1]$ is Pickands dependence function, convex, with

$$\max\{\omega, 1 - \omega\} \leq A(\omega) \leq 1, \forall \omega \in [0, 1].$$

Exemple4

$A(\omega) = 1$ yields the independent copula, C^\perp .

What do we have in dimension 2 ?

Exemple5

$\phi(t) = [-\log(t)]^\theta$ yields Gumbel copula C_θ .

$A(\omega) = [\omega^\theta + (1 - \omega)^\theta]^{1/\theta}$ yields Gumbel copula C_θ .

Definition 14

C is an Archimax copula (from Capéerà, Fougères & Genest (JMVA, 2000)) if

$$C = C_{\phi,A}$$

$$C_{\phi,A}(u, v) = \phi^{-1} \left[[\phi(u) + \phi(v)] A \left(\frac{\phi(u)}{\phi(u) + \phi(v)} \right) \right]$$

Note that there is a frailty type construction, see C. (K, 2006) : given Θ , \mathbf{X} has (survival) copula C_A , Θ has Laplace transform ϕ^{-1} .

Note that $C_{\phi,A}$ is the distorted version of copula C_A .

What do we have in dimension $d \geq 3$?

Definition 15

C is an Archimax copula (from C., Fougères, Genest & Nešlehová (JMVA, 2014)) if

$$C = C_{\phi, \ell}$$

$$C_{\phi, \ell}(u_1, \dots, u_d) = \phi^{-1} [\ell(\phi(u_1) + \dots + \phi(u_d))]$$

This function *is* a copula function.

Stochastic representation of Archimax copulas

Theorem 3

$C_{\phi,\ell}$ is the survival copula of $\mathbf{X} = \mathbf{T}/\Theta$ where Θ has Laplace transform ϕ^{-1} , independent of random vector \mathbf{T} satisfying

$$\mathbb{P}(\mathbf{T} > \mathbf{t}) = \exp[-\ell(\mathbf{t})] = C_{\ell}(e^{-\mathbf{t}}).$$

(see also Li (JMVA, 2009) and Marshall & Olkin (JASA, 1988)).

Limiting behavior of Archimax copulas

One can wonder what would be the max-domain of attraction of that copula ?

$$C_{\phi,\ell} \in \text{MDA}(C_{\ell^*})$$

If $\psi = \phi^{-1}$ is such that $\psi(1-s)$ is regularly varying at 0 with index $\theta \in [1, +\infty]$, then $C_{\phi,\ell}$ belongs to the max domain of attraction of

$$C_{\ell^*}(u_1, \dots, u_d) = \exp \left[-\ell^{\frac{1}{\theta}} (|\log(u_1)|^\theta, \dots, |\log(u_d)|^\theta) \right]$$

(see also C. & Segers (JMVA, 2009) and Larsson & Nešlehová (AAP, 2011) in the case of Archimedean copulas).

forthcoming book (April 2014),
Computational Actuarial Science with R

for additional information

<http://freakonometrics.hypotheses.org/>

